

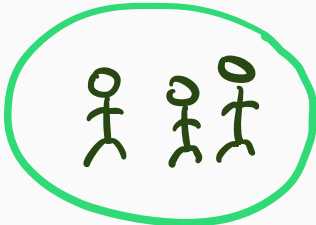
Statistical-Computational Tradeoffs in Mixed Sparse Linear Regression

Gabriel Arpino, Ramji Venkataramanan

October 3rd, 2023



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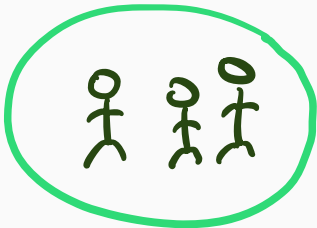
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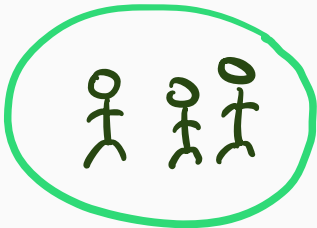
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4	ABQ	Albuquerque	Albuquerque NM	USA	35.04022	-106.60919	
5	ABR	Aberdeen Reg	Aberdeen SD	USA	45.44906	-98.42183	
6	ABY	Southwest Gr	Albany GA	USA	31.53552	-84.19447	
7	ACK	Nantucket M	Nantucket MA	USA	41.25305	-70.06018	
8	ACT	Waco Region	Waco TX	USA	31.61129	-97.23052	
9	ACV	Arcata Airpor	Arcata/Eureka CA	USA	40.97812	-124.10862	
10	ACY	Atlantic City	Atlantic City NJ	USA	39.45758	-74.57717	

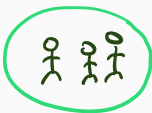
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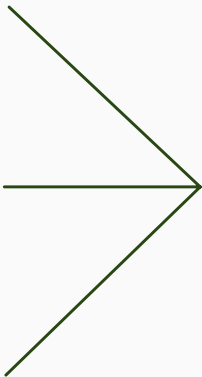
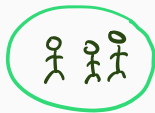
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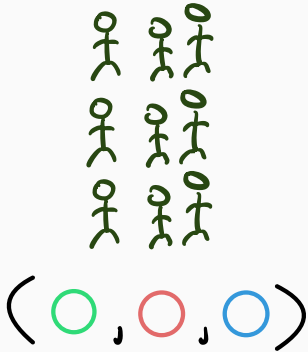


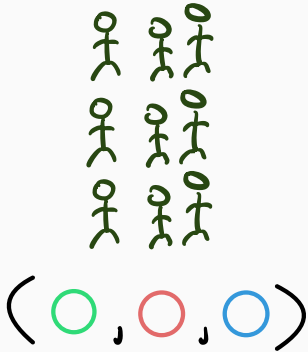


High-Dimensionality

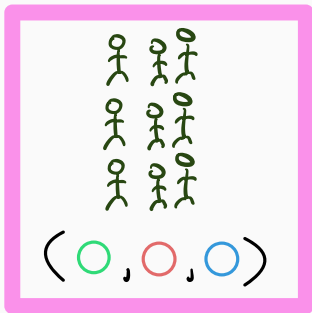


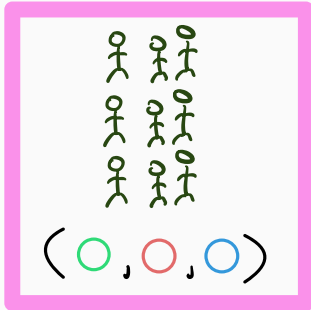






Heterogeneity

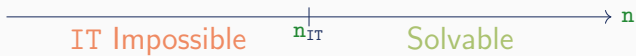


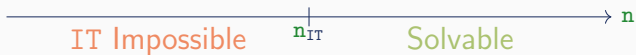


Q: minimal information and computation to recover?

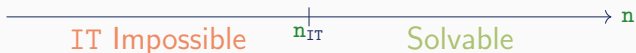


→ n



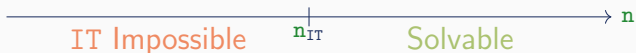


Ex: Communications [S98], Graphical Models [SW09], Regression [MM18], ...

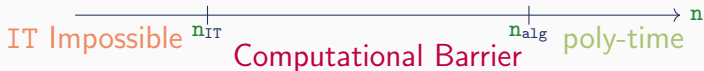


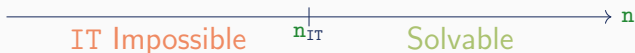
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Ex: Sparse PCA [BR13], Group Testing [CGHWZ22], Sparse Linear Regression [GZ22], ...

Failure of efficient families: MCMC, AMP, Stat Query
[BMNW22, GJ19, BBHLS21, ...]

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 $\iff$ poly-time algorithms

Mixed Sparse Linear Regression (MSLR)

Bad News

Good News

Proof

Mixed Linear Regression

$$y_i = \begin{cases} \langle \mathbf{x}_i, \beta_1 \rangle + w_i, & \text{with indep. prob. } \phi \\ \langle \mathbf{x}_i, \beta_2 \rangle + w_i, & \text{with indep. prob. } 1 - \phi \end{cases}$$

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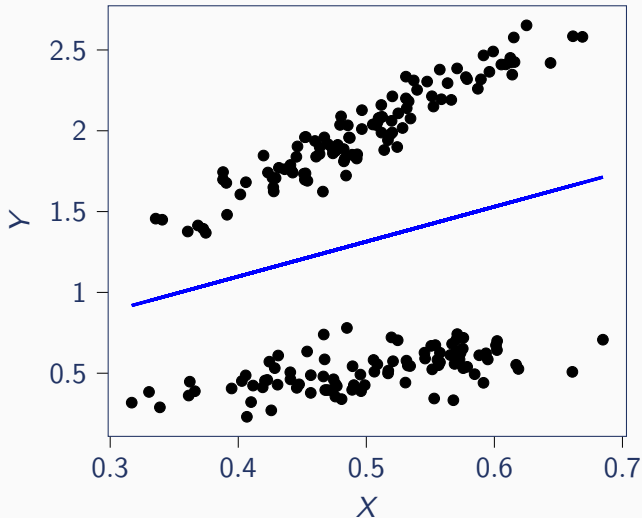
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Goal: Recover β_1, β_2 w.h.p as $p \rightarrow \infty$, given *only* $(y_i, \mathbf{x}_i)_{i \in [n]}$



Applications: *Health Care* [IHHM'22], *Music Perception* [VT'02],
Market Segmentation [WK'00], ...

Model Assumptions

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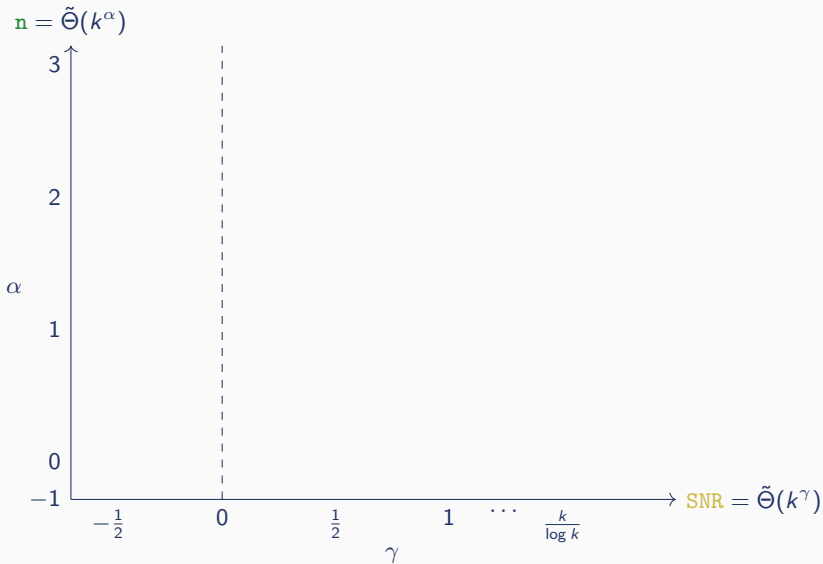
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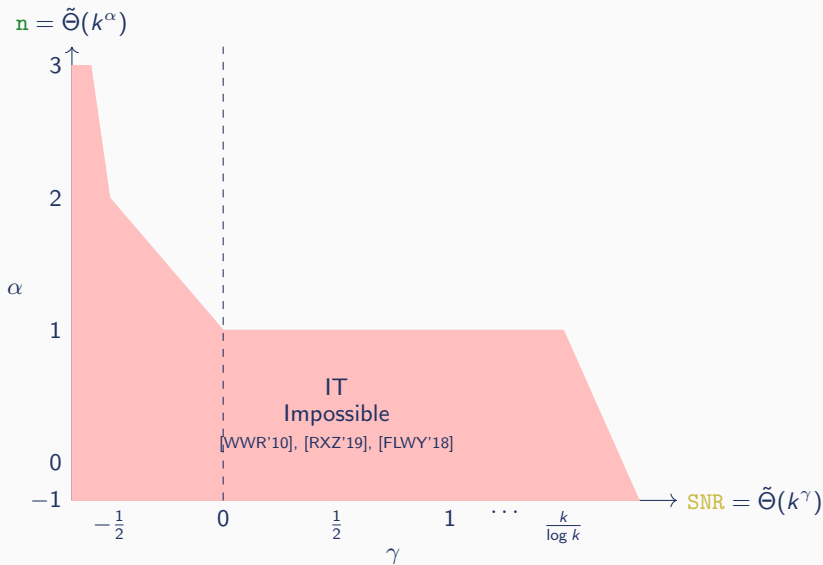
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Linear Regression needs just $\mathfrak{n} = \tilde{\Omega}(k)$ [RXZ'19], [GZ'22]

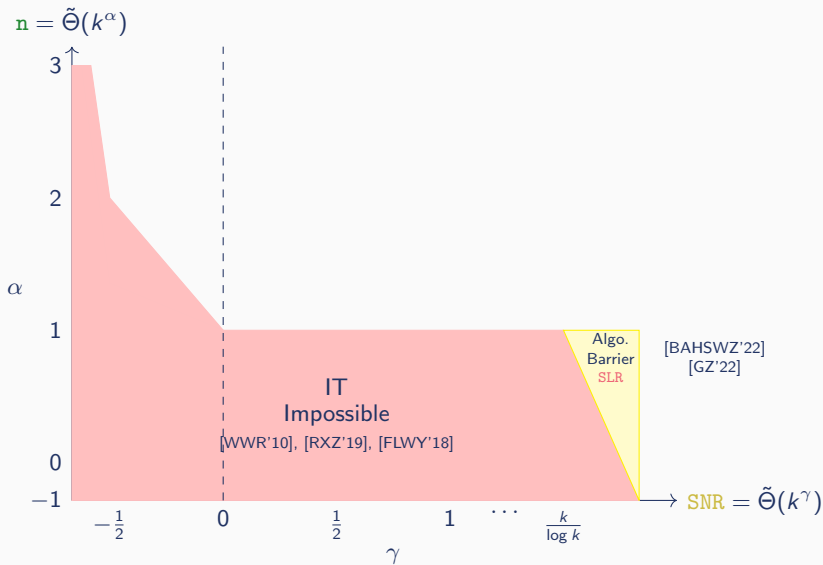
Mixed Sparse Linear Regression



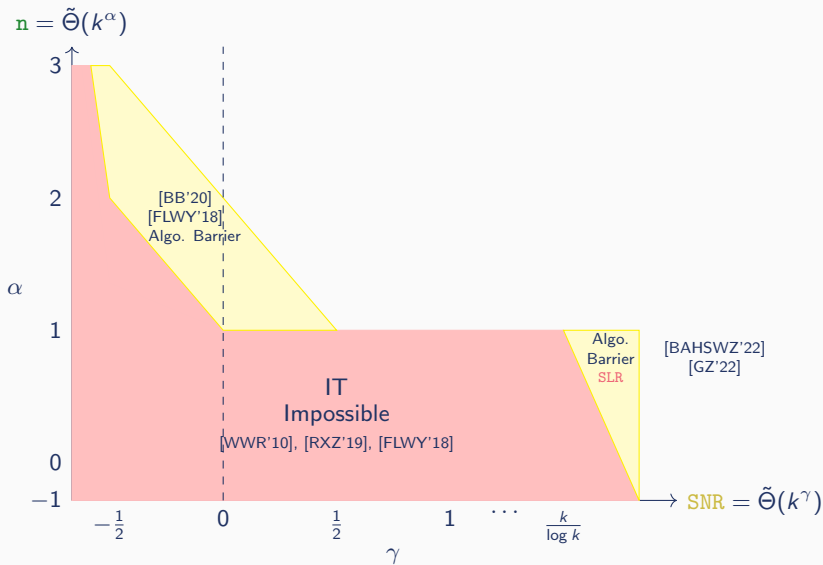
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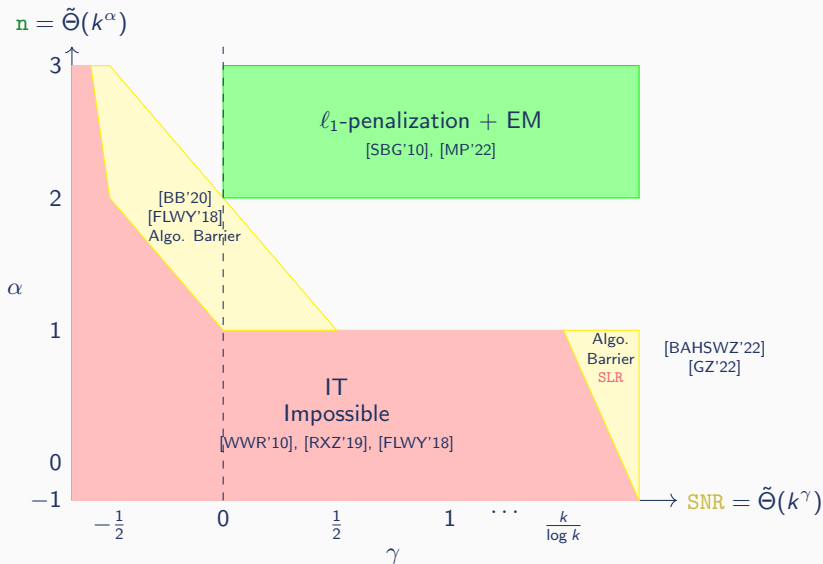
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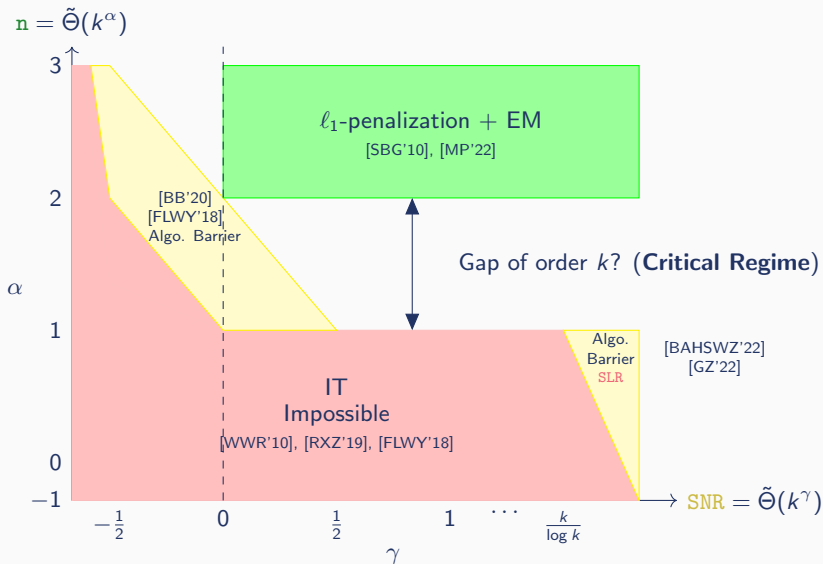
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- $\boldsymbol{\beta}_1 = -\boldsymbol{\beta}_2$, $\phi = 1/2$: **Symmetric Mixed Linear Regression**

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Main Results: Bad News

Theorem. For $k = o(\sqrt{p})$ and $k \ll \underline{n} \ll k^2$, analytic polynomials of degree less than $\frac{k^2}{\underline{n}}$ cannot solve a hypothesis testing variant of Symmetric Mixed Linear Regression.

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Low-Degree Conjecture [H18] + Reduction to Recovery



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Corollary. For $k = o(\sqrt{p})$ and $k \ll \mathfrak{n} \ll k^2$, any randomized algorithm requires runtime $\exp\left(\tilde{\Omega}\left(\frac{k^2}{\mathfrak{n}}\right)\right)$ to exactly recover β_1, β_2 in Symmetric Mixed Linear Regression.

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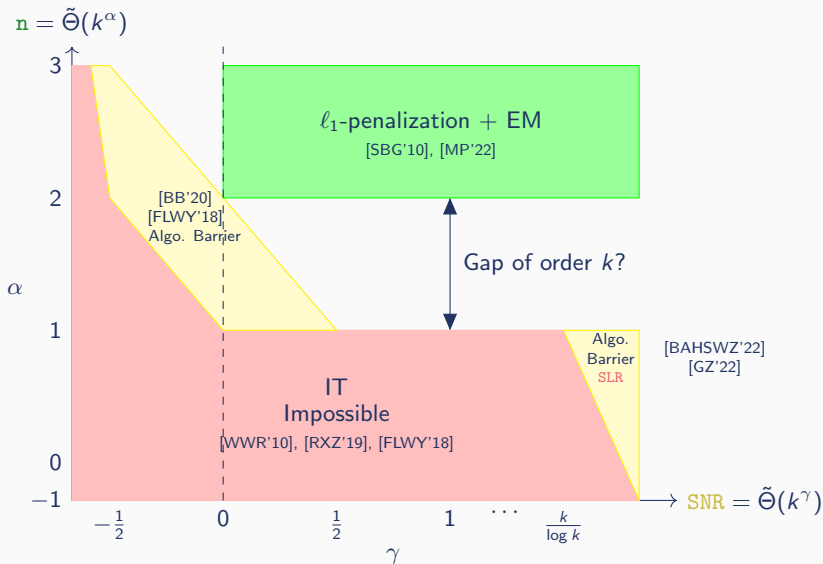
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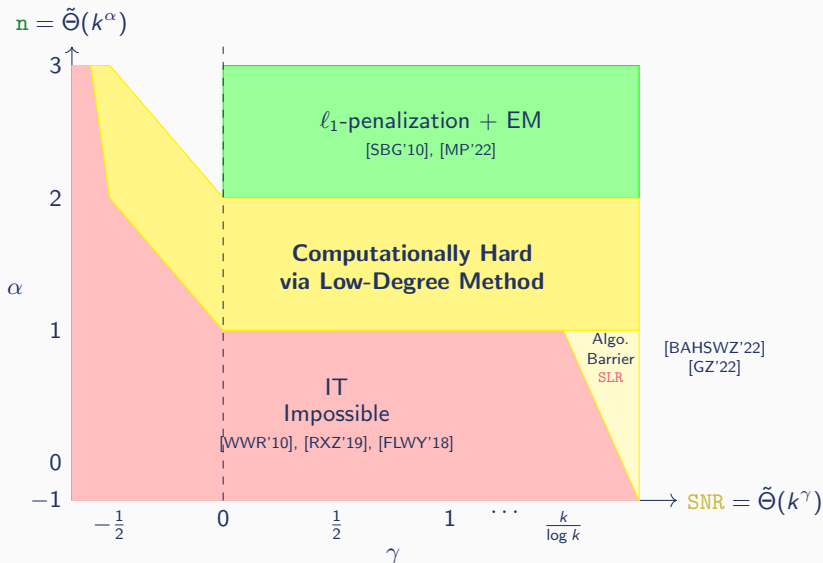
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This runtime is *super-polynomial* in the input!

Mixed Sparse Linear Regression



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Smooth Tradeoff

Corollary. For $k = o(\sqrt{p})$ and $k \ll \mathbf{n} \ll k^2$,

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Recover k -sparse $\mathbf{x}$ from $y_i \stackrel{i.i.d}{\sim} \mathcal{N}(\mathbf{0}, \mathbf{I}_p + \beta \mathbf{x} \mathbf{x}^T)$, $\frac{n}{p} = \Theta(1)$

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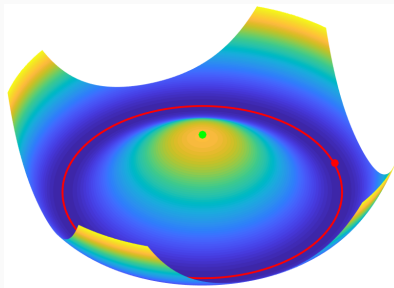
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$$\text{Symmetry} : \boldsymbol{\beta} \mapsto \pm \boldsymbol{\beta}$$

Symmetric Mixed Linear Regression

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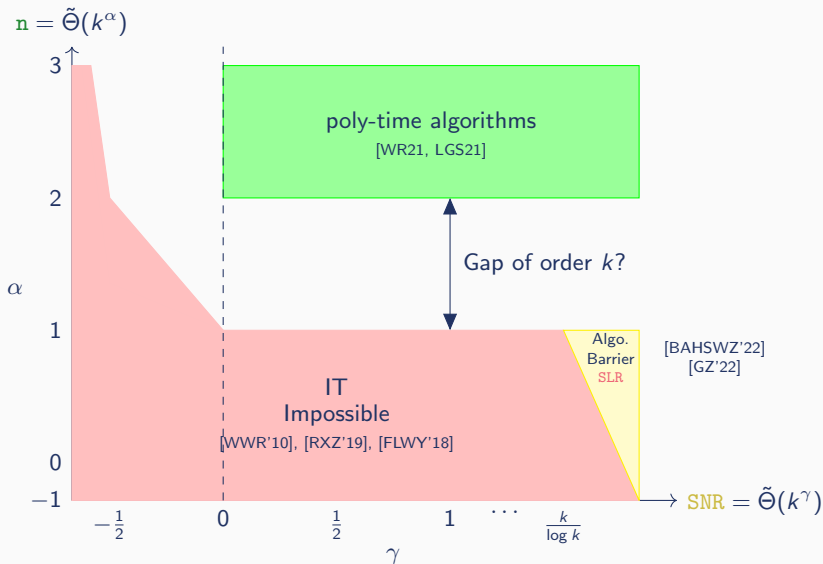
Symmetry : $\boldsymbol{\beta} \mapsto \pm \boldsymbol{\beta}$



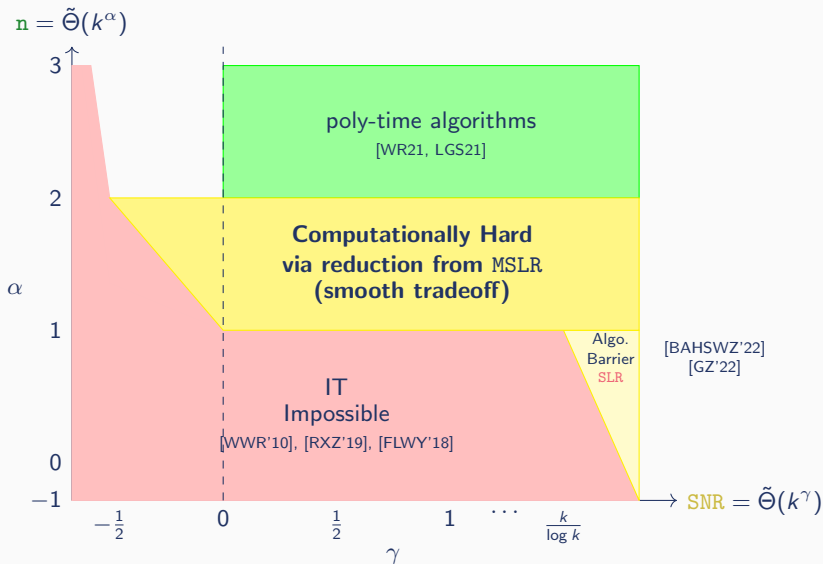
Sparse Phase Retrieval

$$y_i = |\langle \mathbf{x}_i, \boldsymbol{\beta} \rangle| + w_i$$

Sparse Phase Retrieval



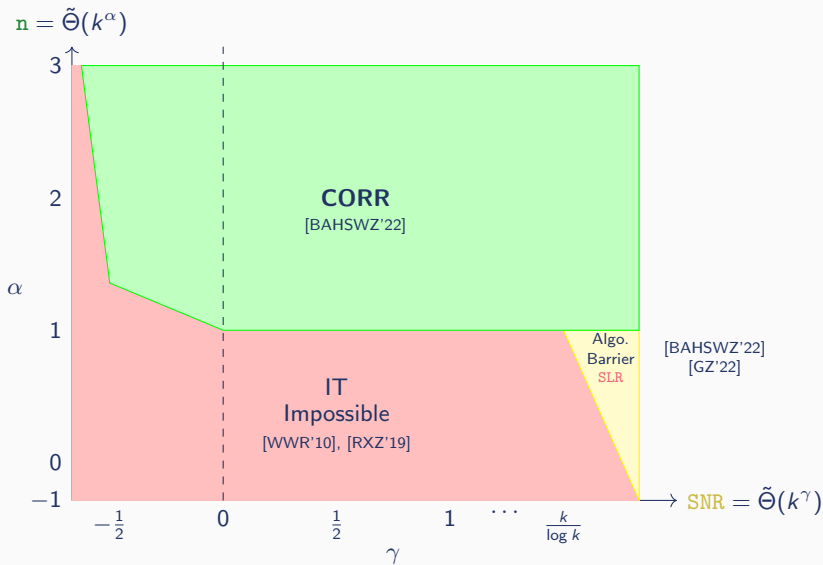
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Theorem. Provided $\mathbf{n} = \tilde{\Omega}(k)$, the **CORR** algorithm recovers the joint support of β_1, β_2 w.h.p.

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Translates to full recovery when combined with dense algorithms
[YCS14, CYC14]

Proof Idea

The Low-Degree Method [HS17]

Hypothesis Test between two distributions with $o(1)$ error:

- Null model (pure noise) $\mathbf{z} \sim \mathbb{Q}_p$
- Planted model (with “signal”) $\mathbf{z} \sim \mathbb{P}_p$

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Low-Deg. Conjecture. [H18, CGHWZ22] If $\|L^{\leq D}\| = O(1)$ for some $D = \omega(\log p)$, then any algorithm distinguishing $\mathbb{P}$ from $\mathbb{Q}$ with $o(1)$ error requires runtime $\exp(\tilde{\Omega}(D))$.

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Proposition. [KWB19] The unique solution f^* to the optimization problem

$$\max_{f \text{ deg } D} \frac{\mathbb{E}_{\mathbf{z} \sim \mathbb{P}} f(\mathbf{z})}{\mathbb{E}_{\mathbf{z} \sim \mathbb{Q}} [f(\mathbf{z})^2]}$$

is the normalized LDLR $f^* = L^{\leq D} / \|L^{\leq D}\|$. The value is $\|L^{\leq D}\|$.

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$$y_i = \begin{cases} \langle \mathbf{x}_i, \boldsymbol{\beta}_1 \rangle + w_i, & \text{with indep. prob. } \phi \\ \langle \mathbf{x}_i, \boldsymbol{\beta}_2 \rangle + w_i, & \text{with indep. prob. } 1 - \phi \end{cases}$$

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Detection Variant

$$y_i = \begin{cases} \langle \mathbf{x}_i, \boldsymbol{\beta}_1 \rangle + w_i, & \text{with indep. prob. } \phi \\ \langle \mathbf{x}_i, \boldsymbol{\beta}_2 \rangle + w_i, & \text{with indep. prob. } 1 - \phi \end{cases}$$

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Goal: Test whether $(y_i, \mathbf{x}_i)_{i \in [n]}$ came from $\mathbb{P}$ or from $\mathbb{Q}$, with $\mathbf{o}(1)$ error as $p \rightarrow \infty$.

$$\mathbf{y} = \mathbf{X}\beta_1 \odot \mathbf{z} + \mathbf{X}\beta_2 \odot (1 - \mathbf{z}) + \mathbf{w}, \quad z_i \stackrel{i.i.d}{\sim} \text{Bernoulli}(\phi)$$

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$\mathbf{X}_j$: column, $\mathbf{x}_i$: row

$$\boldsymbol{\alpha} \in \mathbb{R}^{\mathbf{n} \times (p+1)} : |\boldsymbol{\alpha}| = \sum_{i,j} \alpha_{i,j}, \boldsymbol{\alpha}! = \prod_{i,j} \alpha_{i,j}!$$

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orthonormal basis for poly. deg D w.r.t $\mathbb{Q}$

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orthonormal basis for poly. deg D w.r.t $\mathbb{Q}$

$$\frac{1}{\sqrt{\alpha!}} H_{\alpha}(\mathbf{u}) = \frac{1}{\sqrt{\alpha!}} \prod_{i=1}^n \prod_{j=1}^{p+1} H_{\alpha_{i,j}}(u_{i,j}), \quad \alpha, \mathbf{u} \in \mathbb{R}^{n \times (p+1)}$$

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$$\mathbb{E}_{\mathbf{P}} \left[\prod_{i=1}^{\overset{\text{g}}{\mathbf{n}}} \underbrace{\left(\prod_{j=1}^p H_{\alpha_{i,j}}(X_{i,j}) \right)}_{\text{GIP}} H_{\alpha_{i,p+1}} \left(\underbrace{\frac{(\mathbf{X}\beta_1 \odot \mathbf{z} + \mathbf{X}\beta_2 \odot (1 - \mathbf{z}) + \mathbf{w})}{\sqrt{\|\beta\|_2^2 + \sigma^2}}}_{\mathcal{N}(0,1)} \right) \right]^2$$

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symm.
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\end{aligned}$$

Proof Idea: Good News

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta}_1 \odot \mathbf{z} + \mathbf{X}\boldsymbol{\beta}_2 \odot (1 - \mathbf{z}), \quad z_i \stackrel{i.i.d}{\sim} \text{Bernoulli}(\phi)$$

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[BAHSWZ22]:

$$\text{CORR}(\mathbf{X}, \mathbf{y}) := \left\{ j \in [p] : \left| \frac{\langle \mathbf{X}_j, \mathbf{y} \rangle}{\|\mathbf{y}\|_2} \right| \geq \sqrt{2(1 + \epsilon/2) \log 2p} \right\}$$

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$$\mathbb{E} [\mathbf{X}_{j,i} | y_i, \beta_1, \mathbf{z}_i = 1] = \frac{y_i \cdot \beta_{1,j}}{\|\beta_1\|_2^2 + \sigma^2}$$

$$\mathbb{E} [u_j \mid \mathbf{y}, \mathbf{z}, \beta_1, \beta_2] = \frac{\|\mathbf{y}_{\{\mathbf{z}=1\}}\|_2^2 \beta_{1,j} + \|\mathbf{y}_{\{\mathbf{z}=0\}}\|_2^2 \beta_{2,j}}{\|\mathbf{y}\|_2 (\|\beta\|_2^2 + \sigma^2)}$$

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$$u_j \approx \frac{\phi \mathbf{n}(\|\beta\|_2^2 + \sigma^2) \beta_{1,j} + (1 - \phi) \mathbf{n}(\|\beta\|_2^2 + \sigma^2) \beta_{2,j}}{\sqrt{\mathbf{n}(\|\beta\|_2^2 + \sigma^2)} (\|\beta\|_2^2 + \sigma^2)}$$

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$$= \sqrt{\frac{\mathbf{n}}{\|\beta\|_2^2 + \sigma^2}} (\phi \beta_{1,j} + (1 - \phi) \beta_{2,j})$$

$$\stackrel{\mathbf{n}=\tilde{\Omega}(k)}{>} \sqrt{2(1 + \epsilon/2) \log 2p}.$$

Summary

- Narrow **symmetric** case: $\mathbf{n} = \tilde{\Omega}(k^2)$
- Broad **non-symmetric** case: $\mathbf{n} = \tilde{\Omega}(k)$

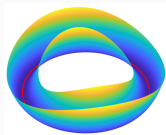
- Sub-exp algorithms for MSLR, Sparse Phase Retrieval

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- Adversarial instances of “robust” inference are mixtures, ex:
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Eigenspace Computation

Compute the principal subspace of a symmetric matrix.

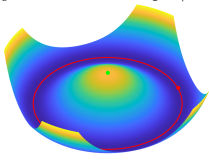


$$\min_{\mathbf{X}} \mathbf{X}^* \mathbf{X} = \mathbf{I} - \frac{1}{2} \text{trace}[\mathbf{X}^* \mathbf{A} \mathbf{X}].$$

Symmetry: $\mathbf{X} \mapsto \mathbf{X} \mathbf{R}$
 $\mathbf{G} = \mathbf{O}(r)$

Generalized Phase Retrieval

Recover a complex vector $\mathbf{x}_o$ from magnitude measurements $\mathbf{y} = |\mathbf{A} \mathbf{x}_o|$.

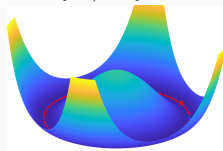


$$\min_{\mathbf{x}} \frac{1}{2} \|\mathbf{y}^2 - |\mathbf{A} \mathbf{x}|^2\|_2^2.$$

Symmetry: $\mathbf{x} \mapsto \mathbf{x} e^{i\phi}$
 $\mathbf{G} = \mathbf{S}^1 \cong \mathbf{O}(2)$

Matrix Recovery

Recover a low-rank matrix $\mathbf{X} = \mathbf{U} \mathbf{V}^*$ from incomplete/corrupted observations



$$\min_{\mathbf{U}, \mathbf{V}} \mathcal{L}(\mathbf{Y} - \mathcal{A}[\mathbf{U} \mathbf{V}^*]) + \rho(\mathbf{U}, \mathbf{V}).$$

Symmetry: $(\mathbf{U}, \mathbf{V}) \mapsto (\mathbf{U} \mathbf{\Gamma}, \mathbf{V} \mathbf{\Gamma}^{-*})$
 $\mathbf{G} = \mathbf{GL}(r)$ or $\mathbf{G} = \mathbf{O}(r)$

[WM22]

Appendix: The Low-Degree Method [HS17]

Goal: *Hypothesis Test* between two distributions with $o(1)$ error:

- Null model (pure noise) $\mathbf{z} \sim \mathbb{Q}_p$
- Planted model (with “signal”) $\mathbf{z} \sim \mathbb{P}_p$

Compute

$$\max_{f \text{ deg } D} \frac{\mathbb{E}_{\mathbf{z} \sim \mathbb{P}}[f(\mathbf{z})]}{\sqrt{\mathbb{E}_{\mathbf{z} \sim \mathbb{Q}}[f(\mathbf{z})^2]}}$$

$\frac{\text{mean in } \mathbb{P}}{\text{fluctuations in } \mathbb{Q}}$

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- Null model (pure noise) $\mathbf{z} \sim \mathbb{Q}_p$
- Planted model (with “signal”) $\mathbf{z} \sim \mathbb{P}_p$

Can we find a low-degree polynomial $f(\mathbf{z})$ that is big for $\mathbf{z} \sim \mathbb{P}$ and small for $\mathbf{z} \sim \mathbb{Q}$?

Compute

$$\max_{f \text{ deg } D} \frac{\mathbb{E}_{\mathbf{z} \sim \mathbb{P}}[f(\mathbf{z})]}{\sqrt{\mathbb{E}_{\mathbf{z} \sim \mathbb{Q}}[f(\mathbf{z})^2]}}$$

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Appendix: The Low-Degree Method [HS17]

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Appendix: The Low-Degree Method [HS17]

$$\begin{aligned} \max_{f \text{ deg } D} \frac{\mathbb{E}_{z \sim \mathbb{P}}[f(z)]}{\sqrt{\mathbb{E}_{z \sim \mathbb{Q}}[f(z)^2]}} \\ = \max_{f \text{ deg } D} \frac{\mathbb{E}_{z \sim \mathbb{Q}}[Lf(z)]}{\sqrt{\mathbb{E}_{z \sim \mathbb{Q}}[f(z)^2]}} \end{aligned}$$

$$L = \frac{d\mathbb{P}}{d\mathbb{Q}}$$

Appendix: The Low-Degree Method [HS17]

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$$\langle f, g \rangle := \mathbb{E}_{z \sim \mathbb{Q}}[f(z)g(z)]$$

Appendix: The Low-Degree Method [HS17]

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$$L = \frac{d\mathbb{P}}{d\mathbb{Q}}$$

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Appendix: The Low-Degree Conjecture

Informal. [H18, CGHWZ22] Assume “sufficiently nice” distributions $\mathbb{P}$, $\mathbb{Q}$. If $\|L^{\leq D}\| = O(1)$ for some $D = \omega(\log p)$, then algorithms require runtime $\exp(\tilde{\Omega}(D))$ to distinguish $\mathbb{P}$ from $\mathbb{Q}$.

Symmetric Mixed Linear Detection

$$y_i = \begin{cases} \langle \mathbf{x}_i, \boldsymbol{\beta} \rangle + w_i, & \text{with indep. prob. } 1/2 \\ -\langle \mathbf{x}_i, \boldsymbol{\beta} \rangle + w_i, & \text{with indep. prob. } 1/2 \end{cases}$$

Under $\mathbb{P}$, observe $\mathbf{x}_i \stackrel{i.i.d}{\sim} \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$, and y_i as above.

Under $\mathbb{Q}$, observe $(y_i, \mathbf{x}_i) \stackrel{i.i.d}{\sim} \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} (\|\boldsymbol{\beta}_1\|_2^2 + \sigma^2) & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_p \end{bmatrix}\right)$.

Goal: Test whether $(y_i, \mathbf{x}_i)_{i \in [n]}$ came from $\mathbb{P}$ or from $\mathbb{Q}$, with vanishing error probability as $p \rightarrow \infty$.

Symmetric Mixed Linear Regression: $\text{SNR} = \infty$

$$y_i = \pm \langle \mathbf{x}_i, \boldsymbol{\beta} \rangle$$

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$$y_i = \pm \langle \mathbf{x}_i, \boldsymbol{\beta} \rangle \stackrel{|\cdot|}{\iff} |y_i| = |\langle \mathbf{x}_i, \boldsymbol{\beta} \rangle|$$

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Corollary. (Conditional on LDC) For $k = o(\sqrt{p})$ and $k \lesssim n \lesssim k^2$, runtime $\exp\left(\frac{k^2}{n}\right)$ is required to solve a detection variant of **Sparse Phase Retrieval**.

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Corollary. (Conditional on LDC) For $k = o(\sqrt{p})$ and $k \lesssim n \lesssim k^2$, runtime $\exp\left(\frac{k^2}{n}\right)$ is required to solve a detection variant of **Sparse Phase Retrieval**.

$\implies$ Provides rigorous evidence for computational hardness of **Sparse Phase Retrieval** when $n = \tilde{o}(k^2)$!